

Rate of convergence in the transfer theorem for min-scheme

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Abstract. In this paper non-uniform estimate of convergence rate in the min-scheme is obtained. Presented results make the estimates, given in [1] and [2], more precise.

Keywords: min-scheme, limit theorem, rate of convergence.

1. Introduction

Let $\{X_j, j \geq 1\}$ be a sequence of independent random variables (r.v.'s) with the common distribution function (d.f.) F . We form r.v.'s

$$W_n = \min(X_1, \dots, X_n), \quad W_{N_n} = \min(X_1, \dots, X_{N_n}),$$

where $\{N_n, n \geq 1\}$ be a sequence of positive integer-valued r.v.'s and independent of $X_j, j \geq 1$, and linear normalized minima

$$\overline{W}_n = d_n^{-1}(W_n - c_n), \quad \overline{W}_{N_n} = d_{N_n}^{-1}(W_{N_n} - c_n), \quad -\infty < c_n < +\infty, \quad d_n > 0.$$

Now, let us denote: $V_n(x) = nF(xd_n + c_n)$.

THEOREM 1 ([1]). *Let*

$$\lim_{n \rightarrow \infty} V_n(x) = v(x) \tag{1}$$

and

$$\lim_{n \rightarrow \infty} P\left(\frac{N_n}{n} \leq x\right) = \lim_{n \rightarrow \infty} A_n(nx) = A(x), \tag{2}$$

then

$$\lim_{n \rightarrow \infty} P(\overline{W}_{N_n} \leq x) = \psi(x), \tag{3}$$

where d.f.

$$\psi(x) = 1 - \int_0^\infty e^{-zv(x)} dA(z).$$

The non-uniform estimate of rate of convergence (deduced from (3)) is presented in this paper. An estimate in the transfer Theorem 1 was obtained in works [2] and [3]. This paper complements the results from Gnedenko [1] and Aksomaitis [2].

2. Main results

A non-uniform estimate of the convergence rate in transfer Theorem 1 is given by the following theorem.

THEOREM 2. *Let conditions (1) and (2) hold and $A(+0) = 0$. Then for any x , satisfying conditions $\frac{V_n(x)}{n} \leq \frac{1}{2}$, the following estimate holds true:*

$$\begin{aligned} \Delta_n(x) &= |P(\overline{W}_{N_n} \leq x) - \psi(x)| \\ &\leq \left(\frac{V_n^2(x)}{n} + |\rho_n(x)|\right) \int_0^\infty z \delta_n^z(x) dA_n(nz) \\ &\quad + v(x) \int_0^\infty |A_n(nz) - A(z)| e^{-zv(x)} dz, \end{aligned} \tag{4}$$

where

$$\begin{aligned} \rho_n(x) &= v_n(x) - v(x), \\ \delta_n(x) &= \max\left(\left(1 - F(xd_n + c_n)\right)^n, e^{-v(x)}\right). \end{aligned}$$

Proof. The complete probability formula implies:

$$\begin{aligned} \Delta_n(x) &= \left| \int_0^\infty (1 - F(xd_n + c_n))^{nz} dA_n(nz) - \int_0^\infty e^{-v(x)z} dA(z) \right| \\ &\leq \left| \int_0^\infty (1 - F(xd_n + c_n))^{nz} dA_n(nz) - Ee^{-\frac{v(x)N_n}{n}} \right| \\ &\quad + \left| Ee^{-\frac{v(x)N_n}{n}} - \int_0^\infty e^{-v(x)z} dA(z) \right| = \Delta_n^{(1)}(x) + \Delta_n^{(2)}(x). \end{aligned} \tag{5}$$

We know that ([2])

$$|u^\alpha - v^\alpha| \leq \alpha (\max(u, v))^\alpha |\ln u - \ln v|, \quad 0 < u, v \leq 1,$$

and

$$\left(1 - \frac{v_n(x)}{n}\right)^n \leq e^{-v_n(x)}, \quad 0 \leq v_n(x) \leq n.$$

These inequalities lead to

$$\begin{aligned} \left| \left(1 - \frac{v_n(x)}{n}\right)^{nz} - e^{-v(x)z} \right| &\leq z \delta_n^z(x) \left| \ln \left(1 - \frac{v_n(x)}{n}\right)^n + v(x) \right| \\ &\leq z \delta_n^z(x) \left(\left| n \ln \left(1 - \frac{v_n(x)}{n}\right) + v_n(x) \right| + |v_n(x) - v(x)| \right). \end{aligned}$$

Using the inequality

$$|\ln(1 - t) + t| \leq t^2, \quad |t| \leq \frac{1}{2},$$

we obtain

$$\left| \left(1 - \frac{v_n(x)}{n}\right)^{nz} - e^{-v(x)z} \right| \leq z\delta_n^z \left(\frac{v_n^z(x)}{n} + |\rho_n(x)| \right). \tag{6}$$

From (5) and (6) it follows that

$$\Delta_n^{(1)}(x) \leq \left(\frac{v_n^z(x)}{n} + |\rho_n(x)| \right) \int_0^\infty z\delta_n^z(x) dA_n(nz). \tag{7}$$

Also

$$\begin{aligned} \Delta_n^{(2)}(x) &= \left| \int_0^\infty e^{-v(x)z} dA_n(nz) - \int_0^\infty e^{-v(x)z} dA(z) \right| \\ &= \left| \int_0^\infty e^{-v(x)z} d(A_n(nz) - A(z)) \right|. \end{aligned}$$

Integrating by parts, we get

$$\begin{aligned} \Delta_n^{(2)}(x) &= \left| \int_0^\infty (A_n(nz) - A(z)) de^{-v(x)z} \right| \\ &\leq v(x) \int_0^\infty |A_n(nz) - A(z)| e^{-v(x)z} dz. \end{aligned} \tag{8}$$

The proof of the theorem follows from (5), (7) and (8).

PROPOSITION 1. *If $EN_n < \infty$, then*

$$\Delta_n^{(1)}(x) \leq \left(\frac{V_n^2(x)}{n} + |\rho_n(x)| \right) \frac{EN_n}{n}.$$

PROPOSITION 2. *If $\rho_n(x) \geq 0$, then*

$$\int_0^\infty z\delta_n^z(x) dA_n(nz) \leq \int_0^\infty ze^{-v(x)z} dA_n(nz).$$

3. Examples

Example 1. Let $P(N_n = k) = \frac{1}{n}, k = \overline{1, n}$.

We get $\lim_{n \rightarrow \infty} A_n(nz) = z, 0 \leq z \leq 1, |A_n(x) - A(x)| \leq \frac{1}{n}, \frac{EN_n}{n} = \frac{n+1}{2n}$.

$$\begin{aligned} \Delta_n^{(1)}(x) &\leq \left(\frac{V_n^2(x)}{n} + |\rho_n(x)| \right) \frac{n+1}{2n}. \\ \Delta_n^{(2)}(x) &\leq \frac{V(x)}{n} \int_0^1 e^{-zv(x)} dz \leq \frac{V(x)}{n}. \end{aligned}$$

For instance, $F(x) = e^x$, $x \leq 0$, then $V_n(x) = ne^{(x-\ln n)} = e^x$, $\rho_n(x) = 0$, and $\Delta_n(x) \leq \frac{e^x}{n} \left(1 + \frac{e^x(n+1)}{2n}\right)$, provided $2e^x \leq n$.

Example 2. Let $P(N_n = k) = \frac{2k-1}{n^2}$, $k = \overline{1, n}$.

We get $\lim_{n \rightarrow \infty} A_n(nz) = A(z) = z^2$, $0 \leq z \leq 1$,

$$|A_n(x) - A(x)| \leq \frac{2nx + 1}{n^2} \leq \frac{2n + 1}{n^2} \leq \frac{3}{n} \cdot \frac{EN_n}{n} = \frac{(n+1)(4n+5)}{6n^2}.$$

$$\Delta_n^{(1)}(x) \leq \left(\frac{V_n^2(x)}{n} + |\rho_n(x)| \right) \frac{(n+1)(4n+5)}{6n^2}.$$

$$\Delta_n^{(2)}(x) \leq \frac{3V(x)}{n}.$$

If $F(x) = e^x$, $x \leq 0$, then $\Delta_n(x) \leq \frac{e^x}{n} \left(\frac{e^x(n+1)(4n+5)}{6n^2} + 3 \right)$.

References

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2. A. Aksomaitis, The nonuniform estimation of the convergence rate in the transfer theorem for extremal values, *Liet. mat. rink.*, **27**, 219–223 (1987).
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REZIUMĖ

A. Aksomaitis. Konvergavimo greičio įvertis minimumų schemos perkėlimo teoremoje

Darbe pateikiamas konvergavimo greičio įvertis stochastinių minimumų schemeje. Rezultatai patikslina Gnedenkų [1] ir A. Aksomaičio [2] gautus rezultatus.

Raktiniai žodžiai: extremes, limit theorem, rate of convergence.